## SIMON FRASER UNIVERSITY School of Engineering Science

### **ENSC 428 Data Communications**

#### **Assignment 5**

### Semester 01-1

### 1. Linear Equalizer

(a) Because three data bits are represented in each matched filter samples, we first have to determine which one we are trying to detect. Recall that the impulse response is

$$x_0 = 0.1$$
  $x_1 = 1$   $x_2 = -0.3$   
so that the samples are  $y_k = \sum_{i=0}^{2} a_{k-i} \cdot x_i + v_k$ 

The strongest sample in x(k) is  $x_1$ , and it exceeds the sum of absolute values of the other samples, so we'll use it to pick off the data. That is, we use  $y_k$  to determine  $a_{k-1}$ .

To determine BER, note that the noise-free value of the samples (i.e., the expected values over the noise ensemble) depends on the data, so there are 8 possible values. By symmetry, we can concern ourselves only with  $a_{k-1}=1$ , reducing the noise-free values to the following:

$$\begin{bmatrix} -1 & 1 & -1 & 1.2 \\ -1 & 1 & 1 & 0.6 \\ 1 & 1 & -1 & 1.4 \\ 1 & 1 & 1 & 0.8 \end{bmatrix}$$
 where  $y_k = 0.1 \cdot a_k + a_{k-1} - 0.3 \cdot y_{k-2}$  (no noise)

The BER depends on the data pattern. Averaging over those patterns, we have

$$P_{b} = \frac{1}{4} \cdot \left( Q\left(\frac{1.2}{\sigma_{v}}\right) + Q\left(\frac{0.6}{\sigma_{v}}\right) + Q\left(\frac{1.4}{\sigma_{v}}\right) + Q\left(\frac{0.8}{\sigma_{v}}\right) \right)$$

Although the average sample value is 1, just as it would be without ISI, the Q function is strongly nonlinear, and the weak values do more harm than the strong values do good. So we lose performance if there is ISI.

(b) One remedy is a simple linear equalizer; i.e., an FIR filter. With 5 coefficients, its impulse response is  $w_j$ , j=0..4 and its output at time k is

$$\sum_{j=0}^{4} \mathbf{w}_{k} \cdot \mathbf{y}_{k-j} = \mathbf{w}^{T} \cdot \mathbf{y} \qquad \text{where} \qquad \mathbf{w} = \begin{bmatrix} \mathbf{w}_{0} \\ \mathbf{w}_{1} \\ \mathbf{w}_{2} \\ \mathbf{w}_{3} \\ \mathbf{w}_{4} \end{bmatrix} \qquad \mathbf{y} = \begin{bmatrix} \mathbf{y}_{k} \\ \mathbf{y}_{k-1} \\ \mathbf{y}_{k-2} \\ \mathbf{y}_{k-3} \\ \mathbf{y}_{k-4} \end{bmatrix}$$

Now we need a compact expression for the components of **y**. Since the oldes component,  $y_{k-4}$ , depends on  $a_{k-4}$  to  $a_{k-6}$ , **y** depends on 7 successive data values.

$$\begin{bmatrix} y_k \\ y_{k-1} \\ y_{k-2} \\ y_{k-3} \\ y_{k-4} \end{bmatrix} = \begin{bmatrix} x_0 & x_1 & x_2 & 0 & 0 & 0 & 0 \\ 0 & x_0 & x_1 & x_2 & 0 & 0 & 0 \\ 0 & 0 & x_0 & x_1 & x_2 & 0 & 0 \\ 0 & 0 & 0 & x_0 & x_1 & x_2 & 0 & 0 \\ 0 & 0 & 0 & x_0 & x_1 & x_2 & 0 \\ 0 & 0 & 0 & 0 & x_0 & x_1 & x_2 \end{bmatrix} \cdot \begin{bmatrix} a_k \\ a_{k-1} \\ a_{k-2} \\ a_{k-3} \\ a_{k-4} \\ a_{k-5} \\ a_{k-6} \end{bmatrix} + \begin{bmatrix} v_k \\ v_{k-1} \\ v_{k-2} \\ v_{k-3} \\ v_{k-4} \end{bmatrix}$$
 or  $\mathbf{y} = \mathbf{X} \cdot \mathbf{a} + \mathbf{m}$ 

Which data symbol will we try to estimate with this equalizer? Since  $y_{k-2}$  is in the centre, following the same logic as in part (a) says we should estimate  $a_{k-3}$ . No surprise - it's in the centre of the **a** vector.

By now, you should not need a rederivation of the normal equations, but I'll throw it in, anyway. You form the estimate

$$\mathbf{a'}_{k-3} = \mathbf{w}^{\mathrm{T}} \cdot \mathbf{y}$$
 with error  $\mathbf{e} = \mathbf{a'}_{k-3} - \mathbf{a}_{k-3} = \mathbf{w}^{\mathrm{T}} \cdot \mathbf{y} - \mathbf{a}_{k-3}$ 

You want to minimize, by choice of  $\mathbf{w}$ , the error variance

$$\sigma_e^2 = E(e^2)$$

Taking the gradient and setting it to (vector)  $\mathbf{0}$  gives

$$\tilde{\mathbf{N}}_{\mathbf{w}} \left( \sigma_{\mathbf{e}}^{2} \right) = E \left( 2 \cdot \tilde{\mathbf{N}}_{\mathbf{w}}(\mathbf{e}) \cdot \mathbf{e} \right) = 2 \cdot E \left( \mathbf{y} \cdot \mathbf{e} \right) = 2 \cdot E \left( \mathbf{y} \cdot \mathbf{y}^{T} \cdot \mathbf{w} - \mathbf{y} \cdot \mathbf{a}_{k-3} \right) = 0$$
  
Using  $\mathbf{R}_{\mathbf{y}} = E \left( \mathbf{y} \cdot \mathbf{y}^{T} \right)$   $\mathbf{p} = E \left( \mathbf{y} \cdot \mathbf{a}_{k-3} \right)$  we have the normal equations  $\mathbf{R}_{\mathbf{y}} \cdot \mathbf{w} = \mathbf{p}$ 

Assuming i.i.d. data and noise samples uncorrelated with each other or the data, we determine

$$\mathbf{R}_{\mathbf{y}} = \mathbf{E} \left[ (\mathbf{X} \cdot \mathbf{a} + \mathbf{n}) \cdot \left( \mathbf{a}^{\mathrm{T}} \cdot \mathbf{X}^{\mathrm{T}} + \mathbf{n}^{\mathrm{T}} \right) \right] = \mathbf{X} \cdot \mathbf{X}^{\mathrm{T}} + \sigma_{v}^{2} \cdot \mathbf{I}_{5}$$
(5x5 identity matrix)

Expansion gives a bulky 5x5 matrix, of which the first 3 columns are

$$\begin{bmatrix} \left(x_{0}\right)^{2} + \left(x_{1}\right)^{2} + \left(x_{2}\right)^{2} + \sigma_{\nu}^{2} & x_{1} \cdot x_{0} + x_{2} \cdot x_{1} & x_{2} \cdot x_{0} \\ x_{1} \cdot x_{0} + x_{2} \cdot x_{1} & \left(x_{0}\right)^{2} + \left(x_{1}\right)^{2} + \left(x_{2}\right)^{2} + \sigma_{\nu}^{2} & x_{1} \cdot x_{0} + x_{2} \cdot x_{1} \\ x_{2} \cdot x_{0} & x_{1} \cdot x_{0} + x_{2} \cdot x_{1} & \left(x_{0}\right)^{2} + \left(x_{1}\right)^{2} + \left(x_{2}\right)^{2} + \sigma_{\nu}^{2} \\ 0 & x_{2} \cdot x_{0} & x_{1} \cdot x_{0} + x_{2} \cdot x_{1} \\ 0 & 0 & x_{2} \cdot x_{0} \\ \end{bmatrix}$$

and the last 2 are

$$\begin{bmatrix} 0 & 0 \\ x_{2} \cdot x_{0} & 0 \\ x_{1} \cdot x_{0} + x_{2} \cdot x_{1} & x_{2} \cdot x_{0} \\ (x_{0})^{2} + (x_{1})^{2} + (x_{2})^{2} + \sigma_{\nu}^{2} & x_{1} \cdot x_{0} + x_{2} \cdot x_{1} \\ x_{1} \cdot x_{0} + x_{2} \cdot x_{1} & (x_{0})^{2} + (x_{1})^{2} + (x_{2})^{2} + \sigma_{\nu}^{2} \end{bmatrix}$$

Substitution of numerical values gives

$$\mathbf{R}_{\mathbf{y}} = \begin{bmatrix} 1.1 + \sigma_{\mathbf{y}}^{2} & -0.2 & -0.03 & 0 & 0 \\ -0.2 & 1.1 + \sigma_{\mathbf{y}}^{2} & -0.2 & -0.03 & 0 \\ -0.03 & -0.2 & 1.1 + \sigma_{\mathbf{y}}^{2} & -0.2 & -0.03 \\ 0 & -0.03 & -0.2 & 1.1 + \sigma_{\mathbf{y}}^{2} & -0.2 \\ 0 & 0 & -0.03 & -0.2 & 1.1 + \sigma_{\mathbf{y}}^{2} \end{bmatrix}$$

Similarly, the **p** vector is given by

$$\mathbf{p} = \mathbf{E} \left( \mathbf{y} \cdot \mathbf{a}_{k-3} \right) = \mathbf{X} \cdot \begin{bmatrix} 0 \\ 0 \\ 0 \\ 1 \\ 0 \\ 0 \\ 0 \end{bmatrix} = \begin{bmatrix} 0 \\ \mathbf{x}_2 \\ \mathbf{x}_1 \\ \mathbf{x}_0 \\ 0 \end{bmatrix} = \begin{bmatrix} 0 \\ -0.3 \\ 1 \\ 0.1 \\ 0 \end{bmatrix}$$

With a value for the noise variance, we could solve  $\mathbf{R}_{\mathbf{y}} \cdot \mathbf{w} = \mathbf{p}$  and obtain the equalizer coefficients that minimize the MSE in estimating  $a_{k-3}$ . To determine the resulting improvement, we would:

\* calculate the new noise variance as

$$\sigma^2 = \sigma_v^2 \cdot \mathbf{w}^T \cdot \mathbf{w}$$

\* calculate the signal amplitude for the  $2^6=64$  combinations of the 6 interfering bits (they interfere a lot less than in part (a)!)

\* average the 64 Q function values.

## 2. Carrier Recovery

There are many ways to handle this problem, and I don't care which way you did it (unless it's hopelessly bad). In these solutions, I have provided:

- \* the decision directed feedforward approach;
- \* the decision directed loop;
- \* the statistically optimum (but computationally intensive) solution.

The first step is to read the data arrays. They are pasted here, and you can scroll up or down with the mouse to see the full array. The lines below convert them to complex format temp2 :=

temp :=

	0	1
0	2.33	0.12
1	-1.08	-1.43
2	1.27	1.81
3	1.37	-0.6
4	-1.83	-0.58
5	0.35	1.67
6	0.65	1.31

matched filter output samples

$$y := temp^{\langle 0 \rangle} + j \cdot temp^{\langle 1 \rangle}$$

 $N_{sym} := length(y) - 1$   $n0 := 0.. N_{sym}$   $n1 := 1.. N_{sym}$ 

A first look at the data suggests that we have a challenge ahead. It seems to be all over the place.



Matched Filter Output Samples

# transmitted data (don't use!)

$$a := temp2^{<0>} + j \cdot temp2^{<1>}$$

0

0

1 2

3

4

5

6

0

0

-1

0

-1

0

1

1

0

-1

0

-1 0

-1

0

No matter which way we approach the problem, some procedures will be useful:

$$\operatorname{decn}(\mathbf{x}) \coloneqq \exp\left[\mathbf{j} \cdot \frac{\pi}{2} \cdot \left(\operatorname{floor}\left(\frac{2}{\pi} \cdot \operatorname{arg}\left(\mathbf{x} \cdot \mathbf{e}^{-\mathbf{j}} \cdot \frac{\pi}{4}\right)\right)\right)\right] \quad \begin{array}{l} \text{This decision function returns} \\ \text{the constellation point nearest} \\ \text{the complex input } x \text{ (there are many ways to write this function).} \end{array}\right]$$

reduce(
$$\theta$$
) := arg $\left(e^{j} \cdot \theta \cdot \overline{\operatorname{decn}\left(e^{j} \cdot \theta\right)}\right)$ 

This one reduces an arbitrary angle  $\theta$  to the range  $-\pi/4$  to  $\pi/4$ , reflecting the four-fold ambiguity.

## The Feedforward Approach

The feedforward approach was outlined in the lecture notes and is sketched below.



The memory of the tracker is the averager. Many ways to average, but the simplest is a recursive averager - a first order lowpass, described variously by

 $v_n = (1 - \alpha) \cdot v_{n-1} + \alpha \cdot u_n \qquad (v_n \text{ is the ouput, } u_n \text{ is the input)}$  $V(z) = \frac{\alpha \cdot z}{z - (1 - \alpha)} \cdot U(z) \qquad h_n = \alpha \cdot (1 - \alpha)^n$ 

dc gain = 1, time constant is about  $1/\alpha$  samples

The tracker operation is captured in the following code. One of the input parameters is the starting value of the averager state. It returns an array with its first column equal to the averager output and second column equal to the b decisions.

track(y, 
$$\alpha$$
, vstart) :=  $\begin{vmatrix} bhat_0 \leftarrow decn(y_0 \cdot \overline{vstart}) \\ v_0 \leftarrow (1 - \alpha) \cdot vstart + \alpha \cdot y_0 \cdot \overline{bhat_0} \\ for \quad n \in 1.. N \quad sym \\ \end{vmatrix}$   
 $\begin{vmatrix} bhat_n \leftarrow decn(y_n \cdot \overline{v_{n-1}}) \\ v_n \leftarrow (1 - \alpha) \cdot v_{n-1} + \alpha \cdot y_n \cdot \overline{bhat_n} \\ augment(v, bhat) \end{vmatrix}$ 

Now we can use it on the samples. Since we don't know anything about the phase to begin, we'll try the following:

vstart := 0.001
$$\alpha := 0.04$$
(try different values and watch the results)est := track(y,  $\alpha$ , vstart)(try different values and watch the results)v := est <0>bhat := est <1>separate the returned valuesahat\_{n1} := bhat\_{n1} · bhat\_{n1-1}differentially decodee :=  $\boxed{(ahat - a) \cdot a}$ find the errors (overhead arrow means "do it to every component")

The results are shown below.



Errors occur in pairs, because an error in *bhat* affects two successive values of *ahat*. The decisions are listed in the Appendix.

Can we improve the tracker with a second pass, starting it where the first pass left off? Let's try:

$$\alpha := 0.01 \qquad \text{est} := \operatorname{track}(y, \alpha, v_{N_{sym}}) \qquad v := \operatorname{est}^{<0>} \qquad \text{bhat} := \operatorname{est}^{<1>}$$

$$\operatorname{ahat}_{n1} := \operatorname{bhat}_{n1} \cdot \overline{\operatorname{bhat}_{n1-1}} \qquad e := \boxed{(\operatorname{ahat} - a) \cdot a}$$



Trajectory of Averager Output

The estimate v is nicely stable.

 $\alpha = 0.01$ 

$$\arg(v_{N_{sym}}) = -0.317$$



Too bad it didn't make any difference to the number of errors - but remember that even perfectly coherent detection produces errors.

## The Tracking Loop

Our next version is a loop. Again, it comes directly from the lecture notes. The sketch below shows that its memory is in the accumulator and the phase error detector (PED) delivers the imaginary part of y' after derotation by the decision *bhat*. The rationale is that the mean PED output should be zero if  $\theta$ *hat* equals the true rotation  $\theta$ , and that  $\theta$ *hat* should be adjusted in proportion to this mean value. Since we don't have the mean value, we'll use the instantaneous value and rely on the loop time constant for the averaging.

In contrast to the feedforward loop, the PLL is strongly nonlinear through the polar to rectangular conversion implicit in the exp(j \*) function. A linearized approximate analysis shows that it operates as a first order averager in  $\theta hat$ , with time constant in samples closely equal to  $(K_cA_y)^{-1}$ , where  $A_y$  is the average value of |y|. It requires more computation than feedforward.



The loop operation is captured in the following code. One of the input parameters is the starting value of the phase estimate  $\theta hat$  stored in the accumulator. The procedure returns an array with its first column equal to the estimated phase and second column equal to the *b* decisions.

$$\begin{split} & \operatorname{loop}(\mathbf{y}, \mathbf{K}_{c}, \theta \operatorname{hatstart}) \coloneqq \quad \mathbf{y}' \leftarrow \mathbf{y}_{0} \cdot \exp(-\mathbf{j} \cdot \theta \operatorname{hatstart}) \\ & \operatorname{bhat}_{0} \leftarrow \operatorname{decn}(\mathbf{y}') \\ & \theta \operatorname{hat}_{0} \leftarrow \theta \operatorname{hatstart} + \mathbf{K}_{c} \cdot \operatorname{Im}\left(\mathbf{y}' \cdot \overline{\operatorname{bhat}_{0}}\right) \\ & \operatorname{for} \quad \mathbf{n} \in 1 \dots \mathbf{N}_{sym} \\ & \quad \left| \begin{array}{c} \mathbf{y}' \leftarrow \mathbf{y}_{n} \cdot \exp\left(-\mathbf{j} \cdot \theta \operatorname{hat}_{n-1}\right) \\ & \operatorname{bhat}_{n} \leftarrow \operatorname{decn}(\mathbf{y}') \\ & \theta \operatorname{hat}_{n} \leftarrow \theta \operatorname{hat}_{n-1} + \mathbf{K}_{c} \cdot \operatorname{Im}\left(\mathbf{y}' \cdot \overline{\operatorname{bhat}_{n}}\right) \\ & \operatorname{augment}(\theta \operatorname{hat}, \operatorname{bhat}) \end{split} \end{split}$$

Now we can use it on the samples. Since we don't know anything about the phase to begin, we'll try the following:

 $\theta$  hatstart := 0 K<sub>c</sub> := 0.1 (try different values and watch the results)



I wonder if a second pass, initialized with the phase estimate inherited from the end of the first pass, would improve things. Reduce the time constant, too. Try it.





The data decisions are listed in the Appendix.

Comparison of the final values of the feedforward estimator and the loop shows they both work.

feedforward loop  $arg(v_{N_{sym}}) = -0.317$   $\theta hat_{N_{sym}} = -0.339$ 

but it's easier to fool the loop if  $\theta$  initially puts the constellation points on decision boundaries.

Finally, the SNR estimate. First, get rid of the data modulation:

$$u := \overline{(y \cdot \overline{bhat})}$$
 overhead arrow means "do it to every component", so *u* is an array of derotated samples, shown below.

This represents the constellation point that is nominally on the positive real axis, rotated by  $\theta$ and buried in noise.

The average value (centroid) is  $\mu_u := mean(u)$   $\mu_u = 1.416 - 0.413i$ 

and its length is  $\sqrt{2 \cdot E_s}$  so  $E_s := \frac{1}{2} \cdot \left( \left| \mu_u \right| \right)^2$   $E_s = 1.088$ 

The real part and the imaginary part of u both have variance  $N_o$ , so

$$N_{o} := \frac{1}{2} \cdot \frac{1}{N_{sym} + 1} \left[ \sum_{i=0}^{N_{sym}} \left( \left| u_{i} - \mu_{u} \right| \right)^{2} \right] \qquad N_{o} = 0.271$$
$$\frac{E_{s}}{N_{o}} = 4.018 \qquad E_{b} := 0.5 \cdot E_{s} \qquad \frac{E_{b}}{N_{o}} = 2.009$$

### The Statistically Optimum Solution

Both feedforward and the loop have the appearance of being *ad hoc*, approximate solutions. They are. Now let's see what the optimum solution is for a block of N symbols. Refer to page Section 7.5 of the lecture notes.

We have a length-*N* vector of samples **r** and a large number of candidate signal vectors  $\mathbf{s}_i$ , *i*=1..*M*, each corresponding to a different vector **b**. Large does mean *large*: for QPSK, *M*=4<sup>*N*</sup>. We have the basic pdf

$$p_{\mathbf{r}|\mathbf{s}_{i},\theta}(\mathbf{r} \mid \mathbf{s}_{i},\theta) = p_{\mathbf{n}}(\mathbf{r} - e^{j\theta}\mathbf{s}_{i}) = \frac{1}{\left(2\pi N_{o}\right)^{N}} \exp\left(-\frac{\left(\mathbf{r} - e^{j\theta}\mathbf{s}_{i}\right)^{H}\left(\mathbf{r} - e^{j\theta}\mathbf{s}_{i}\right)}{2N_{o}}\right)$$

where superscript H denotes Hermitian transpose (conjugate transpose). Expanding the exponent reveals factors that do not affect the decision

$$p_{\mathbf{r}|\mathbf{s}_{i},\theta}(\mathbf{r} \mid \mathbf{s}_{i},\theta) = \frac{1}{\left(2\pi N_{o}\right)^{N}} \exp\left(-\frac{\mathbf{r}^{H}\mathbf{r} + 2NE_{s}}{2N_{o}}\right) \exp\left(\frac{\operatorname{Re}\left[e^{-j\theta}\mathbf{s}_{i}^{H}\mathbf{r}\right]}{N_{o}}\right)$$

since  $\mathbf{s}_{i}^{H} \cdot \mathbf{s}_{i} = 2 \cdot \mathbf{N} \cdot \mathbf{E}_{s}$ 

We don't know  $\theta$ , so we consider it a nuisance parameter, and form the marginal density by integrating it out:

$$p_{\mathbf{r}|\mathbf{s}_{i}}(\mathbf{r} \mid \mathbf{s}_{i}) = \int_{0}^{2\pi} p_{\mathbf{r},\theta|\mathbf{s}_{i}}(\mathbf{r},\theta \mid \mathbf{s}_{i})d\theta = \int_{0}^{2\pi} p_{\mathbf{r}|\mathbf{s}_{i},\theta}(\mathbf{r} \mid \mathbf{s}_{i},\theta) p_{\theta}(\theta)d\theta = \frac{1}{2\pi} \int_{0}^{2\pi} p_{\mathbf{r}|\mathbf{s}_{i},\theta}(\mathbf{r} \mid \mathbf{s}_{i},\theta)d\theta$$

Substituting for the conditional density in the last equality above, ignoring the factors that do not affect the decision and integrating as in Section 7.5, we obtain

$$\hat{i} = \arg\max_{i} \left( I_0 \left( |\mathbf{s}_I^H \mathbf{r}| \right) \right) = \arg\max_{i} \left( |\mathbf{s}_I^H \mathbf{r}|^2 \right)$$

This means that we should correlate the N-sample received vector with every possible transmitted sequence and select the correlation with the largest magnitude (or magnitude squared). This could be a *lot* of work for long sequences. However, for shorter blocks, such as 4 to 6 symbols, the method received a brief flurry of interest as *multisymbol differential detection* in the context of wireless communication.

## **APPENDIX: DATA DECISIONS**

The feedforward and loop structures produce the same sequence of decisions. The array is broken into a first half and a second half so they fit on one page.

ahatfirst := submatrix(ahat, 0, 25, 0, 0) ahatsecond := submatrix(ahat, 26, 50, 0, 0)

