

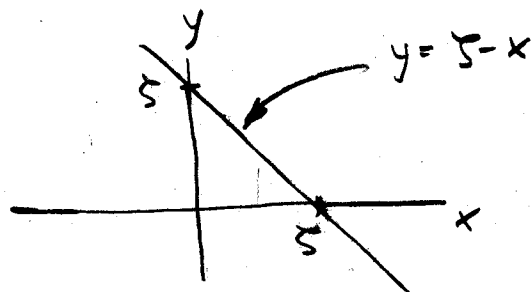
2.3 Sums of Random Variables

- Suppose $z = x + y$ and we have $f_{xy}(x, y)$.
What is $f_z(z)$?

Easy answer through cdf:

$$F_z(z) = \text{Prob}[x + y \leq z]$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{z-x} f_{xy}(x, y) dy dx$$



then $f_z(z) = \frac{d}{dz} F_z(z)$

- If x, y independent (important case), then

$$F_z(z) = \int_{-\infty}^{\infty} \int_{-\infty}^{z-x} f_x(x) f_y(y) dx dy = \int_{-\infty}^{\infty} F_y(z-x) f_x(x) dx$$

$$f_z(z) = \frac{d}{dz} F_z(z) = \int_{-\infty}^{\infty} f_y(z-x) f_x(x) dx$$

- The F transform of a pdf (the "characteristic function") simplifies it:

$$M_z(\omega) = \int_{-\infty}^{\infty} f_z(z) e^{-j\omega z} dz$$

$$M_z(\omega) = M_x(\omega) M_y(\omega)$$

The characteristic function is sometimes called the "moment generating function"

$$\frac{d}{d\omega} M_x(\omega) = \int -jx f_x(x) e^{-j\omega x} dx, \quad M'(\omega)|_{\omega=0} = -j\mu$$

$$\frac{d^2}{d\omega^2} M_x(\omega) = \int -x^2 f_x(x) e^{-j\omega x} dx, \quad M''(\omega)|_{\omega=0} = -\overline{x^2}$$

$$\text{so } \sigma^2 = -M''(0) + M'(\omega)^2$$

The characteristic function of a jointly Gaussian random vector \underline{x} of length N has N frequency variables in $\underline{\omega}$

$$M_x(\underline{\omega}) = E[e^{-j\underline{\omega}^T \underline{x}}] = \int \dots \int e^{-j(\omega_1 x_1 + \dots + \omega_N x_N)} P_{x_1, \dots, x_N}(x_1, \dots, x_N) dx_1 \dots dx_N$$

$$= \exp(-\frac{1}{2} \underline{\omega}^T C \underline{\omega}) \text{ for Gauss, zero mean cov } C$$

- The mean of a sum is the sum of the means.
Always?

• What about the variance of a sum?

$$z = \sum_i x_i \quad m_z = \sum_i m_i$$

$$\sigma_z^2 = \overline{(z - m_z)^2} = \overline{\left(\sum_i x_i - m_i\right)^2} = \overline{\sum_i \sum_j (x_i - m_i)(x_j - m_j)}$$

$$= \sum_i \sum_j \sigma_{ij}^2 = \sum_i \sigma_i^2 + 2 \sum_i \sum_{j \neq i} \sigma_{ij}^2$$

Aside: another way to get this is to pack $\{x_i\}$ into \underline{x} .

Then $z = \underline{e}^T \underline{x}$, where $\underline{e} = \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix}$

and $z - m = \underline{e}^T (\underline{x} - \underline{m}_x)$

$$\overline{(z - m)^2} = \overline{\underline{e}^T (\underline{x} - \underline{m}_x) (\underline{x} - \underline{m}_x)^T \underline{e}} = \underline{e}^T C_x \underline{e}$$

$$= \sum_i \sum_j \sigma_{ij}^2$$

→ variance of a sum equals sum of variances iff
 r.v.s are uncorrelated; i.e. $\sigma_{ij}^2 = 0$ for $i \neq j$
 i.e. C is diagonal

• Add (or average) N i.i.d. variates

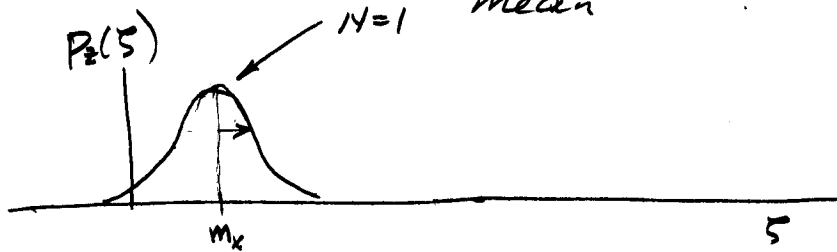
- if $z = \sum_{i=1}^N x_i$

then $m_z = N m_x$ $\sigma_z^2 = N \sigma_x^2$

but it's σ that we compare with m , not σ^2

$\sigma_z = \sqrt{N} \sigma_x$ scatter increases more slowly than mean

for $N=4$?



Note - not necessarily Gaussian to start

- coefficient of variation measures scatter relative to mean

$C = \frac{\sigma}{m}$. For z , $C_z =$

- average: $z = \frac{1}{N} \sum_{i=1}^N x_i$

so $m_z = m_x$ $\sigma_z^2 = \frac{1}{N^2} N \sigma_x^2 = \frac{\sigma_x^2}{N}$ $\sigma_z =$

